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Powerlaw scaling in the internal variability of cumulus cloud size distributions due to subsampling and spatial organization

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ABSTRACT

In this study the spatial structure of cumulus cloud populations is investigated using three-dimensional snapshots from large-domain LES experiments. The aim is to understand and quantify the internal variability in cloud size distributions due to subsampling effects and spatial organization. A set of idealized shallow cumulus cases is selected with varying degrees of spatial organization, including a slowly organizing marine precipitating case and five more quickly organizing diurnal cases over land. A subdomain analysis is applied, yielding cloud number distributions at sample sizes ranging from severely undersampled to nearly complete. A strong powerlaw scaling is found in the relation between cloud number variability and subdomain size, reflecting an inverse linear relation. Scaling subdomain size by cloud size yields a data collapse across timepoints and cases, highlighting the role played by cloud spacing in controlling the stochastic variability. Spatial organization acts on top of this baseline model, by increasing the maximum cloud size and by enhancing the variability in the number of smallest clouds. This reflects that the smaller clouds start to live on top of larger-scale structures, favoring or inhibiting their formation. Compositing all continental cumulus cases suggests the existence of a prototype diurnal time-dependence in the spatial organization. A simple stochastic expression for cloud number variability is proposed, formulated in terms of two dimensionless groups and accounting for both subsampling and spatial organization. The implications of the results for the stochastic and scale-adaptive modeling of convection in the grey zone are briefly discussed.

1. Introduction

Recent advances in supercomputing have introduced a "grey zone" in the representation of cumulus convection in General Circulation Models (GCMs), in which this physical process is getting partially resolved (Wyngaard 2004). Existing parameterization schemes, built on the assumption that cumulus populations are fully sampled in the GCM-gridbox, often lack the scale-awareness and scale-adaptivity to conceptually and practically deal with this situation. How to best achieve this is still an open research question.

Research into the spatial structure of cumulus cloud populations goes back decades (e.g. Cahalan and Joseph 1989; Sengupta et al. 1990; Nair et al. 1998), but has recently intensified due to the arrival of the grey zone problem. A natural way of characterizing the scale-dependence within a cumulus population is the *size distribution* of a cloud field, which has been scientifically established for many cloud regimes using a large variety of instrumentation (e.g. Plank 1969; Raga et al. 1990; Benner and Curry 1998; Zhao and Di Girolamo 2006; Yuan 2011) and finescale cloud-resolving simulation (e.g. Neggers et al. 2003; Rieck et al. 2014; Senf et al. 2018). A characteristic feature of the size density of cumulus cloud number is its shape, which has been described by lognormal, exponential and powerlaw functions. These shapes are argued to reflect the underlying processes of cloud formation, interaction and subcritical percolation (Cohen and Craig 2006; Ding et al. 2014). The size-range covered by the distribution has been found to differ between marine and continental cloud populations (Sakradzija and Hohenegger 2017). Over land the maximum size of the distribution exhibits a distinct diurnal dependence . With many studies relying on fine-scale simulations of cumulus cloud populations, the confrontation of simulated and observed cumulus populations has recently become an important focus point. Both area-covering measurements (Kassianov et al. 2005) and vertically pointing measurements (Lareau et al. 2018) have started to provide a wealth of new information about cumulus cloud populations at high resolutions and frequencies.

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Building on these new results and insights, a new class of convection schemes has recently emerged in which the transport is explicitly formulated in terms of discretized Cloud Size Densities (CSDs, e.g. Wagner and Graf 2010; Park 2014; Neggers 2015). In essence these schemes adopt the spectral approach in convective modeling as first proposed by Arakawa and Schubert (1974). Compared to bulk convection schemes the CSD is a new variable that requires closure, concerning both its functional form and its range. Differences exist among schemes in how the CSD is reconstructed. But a general benefit is that the CSD is in general well observable, allowing such models to be constrained by both existing and new datasets. Furthermore, scale-awareness is by definition present at the foundation of the scheme (Neggers 2015). The reconstructed CSD can be size-filtered, which introduces scale-adaptivity in the parameterized convective transport and clouds (Brast et al. 2018).

A complicating factor is that the CSD is not a global constant and exhibits both external and internal variability. External variability arises from various sources such as large-scale synoptic conditions or surface properties (Rieck et al. 2014). In constrast, internal variability arises due to spatial sampling issues. Given a well defined CSD reflecting a cloud population covering an infinitely large area, a limited spatial sampling leads to stochastic variability (e.g. Nair et al. 1998; Cohen and Craig 2006). This situation applies to convective modeling in the grey zone, in which the gridbox has become too small to include the full population (Dorrestijn et al. 2013; Honnert et al. 2011). In CSDs diagnosed in LES realizations of a relatively small domain this stochastic signal is sometimes present the right tail, where large clouds are barely sampled (Neggers et al. 2003). The spatial organization in a cloud field can add to this stochastic variability, by introducing perturbations in the cloud populations on scales much larger than the boundary layer depth (e.g. Seifert and Heus 2013). Physical-dynamical processes that drive spatial organization include to cold pool formation (e.g. Schlemmer and Hohenegger 2014) and oscillations (Sakradzija et al. 2014; Feingold et al. 2017).

In recent years super-large-domain LES has become computationally feasible (Khairoutdinov and Randall 2006; Satoh et al. 2008; Heinze et al. 2016), and has offered new ways to study the CSD and its variability. The availability of high spatial resolution in combination with a large domain size yields cloud size distributions that can freely form and evolve across a range of scales that is less and less artificially constrained at both ends. The large domain allows sampling the cumulus population more completely, also covering the rarely occurring largest clouds better. With the cumulus clouds still resolved, the ever larger domain of the simulation also allows mesoscale fluctuations to form naturally and start to affect the CSD. The opportunities created by super-large-LES to investigate the behavior of cumulus CSDs and their interaction with mesoscale organization have only recently been started to be explored (e.g. Senf et al. 2018).

The goal of this study is to use large-domain simulations to investigate and quantify the internal variability in a cumulus CSD. We ask how this variability depends on i) the domain size of the analysis and ii) the degree of spatial organization within the cloud population. The aim is to work with domain sizes large enough to approach full sample size when diagnosing the CSDs. Two different convective regimes are investigated. The first is a slowly organizing marine cumulus case based on the RICO field campaign (Rauber et al. 2007). The second regime reflects continental summertime conditions at the Southern Great Plains (SGP) site of the Atmospheric Radiation Measurement program (ARM, Stokes and Schwartz 1994). A subdomain analysis is performed for all fields to quantify the impact of subsampling on the CSD. The difference in organization speed between the two regimes then provides insight on how this process affects the internal variability of shallow cumulus CSDs. TThe implications of the obtained results for our understanding of cumulus populations in general and for stochastic cumulus parameterization in the grey zone in particular will be discussed.

2. Method

a. Large eddy simulations

This study makes use of LES results for two welldefined shallow cumulus regimes, in which the pace of the spatial organization differs considerably. The first, a slowly organizing cloud regime, reflects marine subtropical subsidence conditions as observed during the Rain in Cumulus over the Ocean (RICO) field campaign (Rauber et al. 2007). The spatial organization in this case is related to warm precipitation processes, producing cold pools with low cloud mass that are separated by convergence lines where clouds tend to congregate. The spatial organization in this case is a slow process, taking multiple days. The second, faster organizing cloud regime reflects diurnal cycles of continental shallow cumulus at the ARM SGP site. For this regime five day-long cases are chosen, which are all part of the archive of simulations for the ARM SGP site in 2016 generated by the LES ARM Symbiotic Simulation And Observation Workflow (LASSO, Gustafson et al. 2017b). The cases include 18 May, 19 and 25 June, 16 July and 18 August 2016 (from now on be referred to as the 20160518, 20160619, 20160625, 20160716 and 20160818 cases). These days are selected because they best reflect the prototype view of diurnal cycles of shallow convection, with most daytime cloudiness associated with surface driven convective clouds and with more or less clear-sky conditions before and afterwards.



FIG. 1: Vertically projected cloud masks for the six timepoints during the LES simulation of the RICO case that are considered in this study. The domain size shown is 51.2 km. A gridpoint is masked (black) in case any condensate exists in the vertical column.



FIG. 2: Similar to Fig. 1, but now showing the cloud mask for 6 timepoints during the LASSO 20160518 case. Domain size is 25km.

The LES experiments for these 6 cases have already been described in detail in the literature, and only the details relevant for this study will be summarized here. The RICO composite shallow cumulus case (vanZanten et al. 2011) is simulated using the UCLA-LES code, as described in detail by Heus and Seifert (2013); their simulations are also used in this study. The simulated domain size is $D_x \times D_y \times D_z = 51.2 \times 51.2 \times 4.0$ km. A period of 48 hours is simulated, considered long enough for the spatial organization to take place. Full three-dimensional fields are stored at 8 hour intervals. The LASSO cases have been resimulated with the MESOHH code (van Heerwaarden et al. 2018), using the WRF-based prescribed large-scale forcings and boundary conditions that are part of the LASSO Alpha 2 dataset (Gustafson et al. 2017a). Compared to the standard LASSO simulations the horizontal domain size is extended to 25.6 km in the resimulations in order to maximize the domain-range available for the subdomain analysis of CSDs.

Although two different LES codes are used, the experiment setups share some important aspects. A spatial resolution of $25 \times 25 \times 25$ m is always applied, as is adaptive



FIG. 3: Simulated and observed total cloud cover for the 5 LASSO cases investigated in this study. TSI is an estimate of the total sky cover and is only available during day time, while ARSCL represents a cloud cover as diagnosed over only the lowest 5km. The grey bars denote the period before sunrise and after sunset.

time-stepping. Both the UCLA-LES and the MICRO-HH codes apply the Smagorinsky scheme for subgrid transport of momentum, energy and heat, and both make use of the warm-cloud double-moment microphysics scheme of Seifert and Beheng (2001). Horizontally periodic boundary conditions are applied, as well as a sponge-layer in the top third of the domain to dampen any gravity waves. A prescribed surface temperature is used in both the RICO and LASSO cases, the only difference being the presence of a diurnal cycle in the latter.

The evolution of the simulated cloud population for the RICO case is illustrated in Fig. 1, showing the vertically projected cloud mask at each storage timepoint. The cloud field at the first two time-points still looks more or less homogeneously distributed, but increasingly the cloud field becomes more organized, with large cloud structures appearing after about one day of simulation. Figure 2 shows the same for the LASSO 20160518 case, revealing the distinct diurnal evolution of the cloud population. The spacing between the largest clouds increases during the day, likely reflecting the continuous ongoing deepening of the boundary-layer that is typical of such diurnal cycles (Brown et al. 2002; Zhang et al. 2017). Organization in the cloud field becomes apparent during the final period, with small clouds increasingly clustering around the larger ones, being separated by large areas without any cloud mass. This may be related to the decay of largest clouds at the end of the day, when the surface forcing of the turbulence weakens.

To provide further confidence in the realism of the simulations the LES cloud cover for the five LASSO cases is compared to two ARM observational products. The AR-SCL value-added product (Clothiaux et al. 2000, 2001) combines lidar and radar measurements, while the Total Sky Imager (Kassianov et al. 2005) retrieves cloud cover from wide-angle imagery. The comparison for the five cases is shown in Fig. 3, only showing daytime hours. While some differences exist between the two measurements, they agree reasonably well on the general evolution and amplitude of the cloud cover. The model does reproduce these trends, however the amplitude seems to be a bit underestimated. This slight bias has been revealed in other recent LES evaluation studies at supersites (Schalkwijk et al. 2015; Zhang et al. 2017). In addition, on 18 May and 18 August the LES seems to underestimate cloud cover in the morning, which we speculate is due to either i) high altitude cloudiness or ii) spin-up effects. Despite these minor shortcomings, the results do suggest that the experiment setup captures the prototype diurnal variation in cloudiness present in the observations to a reasonable degree, as well as its modulation due to large-scale forcing. This agreement suggests the simulations are representative of nature, and justifies their use for subsequent cloud studies.

b. Deriving cloud size distributions

Cloud size distributions can be derived in many ways, involving choices concerning what defines a cloud and its size. In this study a simple definition is adopted, to be applicable to both observed and simulated cloud populations, and that allows using past data records as much as possible. The first simplification is to disregard cloud lifecycles, by only analyzing standalone instantaneous threedimensional fields. Often the time-frequency of observations of cloud populations is simply too low to allow detection of cloud life-cycle properties. Also, we hypothesize that the CSD-internal variability should already be detectable in instantaneous snapshots. The second choice is to use the presence of any cloud condensate mass to label a gridbox as cloudy or non-cloudy. This allows the analysis to be used with any LES simulation, no matter how complex its representation of cloud microphysics. The third choice is to define cloud size as the square-root of the projected area fraction of the object defined on the 3D grid. This choice is made to allow comparison to previous CSD studies which often relied on this definition.

In practice the "cusize" algorithm is applied as described by Neggers et al. (2003) to investigate diurnal cycles in the CSD. Given a 3D field of cloud mask, first all objects are identified that consist of neighboring cloudy gridboxes and which are separated by clear air. The size of each object is then established as the square-root of its projected cloud cover. These are then sorted into a histogram



FIG. 4: Cloud number densities \mathcal{N} for a) the RICO case and b) the LASSO 20160518 case, as derived from the fulldomain LES fields. Each line represents a single instantaneous 3D cloud field, with *t* its timepoint in hours. The clouds were sorted in histograms using a constant bin-width of 25 m, matching the LES resolution.

of cloud number N(l), yielding the CSD $\mathcal{N}(l)$ defined as

$$\mathcal{N}(l) = N(l) \, \mathrm{d}l^{-1},\tag{1}$$

with dl the width of the size-bins. As illustration the resulting size densities of cloud number in log-log space for the RICO case and the LASSO 20160518 case are shown in Fig. 4. Each line represents a single instantaneous 3D field. The typically observed behavior that cumulus cloud occurrence sharply decreases with size is also evident in these cases, as expected. The RICO case shows powerlaw scaling in the size-range l < 300m, above which the CSD decreases more rapidly with size. The first cloudy hours in the LASSO case also has this shape, but as time progresses the slope of the CSD becomes more constant across the covered size-range. Note that a higher vertical position in the frame corresponds to a larger number of clouds, due to the use of non-normaliued densities. In the LASSO case, cloud number strongly increases across the spectrum after cloud onset, and decreases again towards the end of the day. What both cases share is the increasing spread on the vertical axis (number density) towards the larger sizes, in the right tail. This spread, from now on referred to as CSD-internal variability or stochastic variability, is the main focus of this study.

c. Subdomain analysis

The subdomain analysis method used by Dorrestijn et al. (2013) to investigate the scale-dependence of turbulent transport is applied here to cloud number N. To

this purpose the full horizontal LES domain $D_x \times D_y$ is first subdivided into small square subdomains of horizontal size $L_0 = 1.6$ km. This basic size was chosen to still contain a sufficient number of gridpoints on a horizontal slice (64²), but on the other hand to still be significantly smaller than the typical neighbor spacing of large cumulus clouds (Joseph and Cahalan 1990). Also, with this value of L_0 the full domain size is exactly covered for both the RICO and LASSO cases.

The next step is to define a set of bigger square subdomains of size L that consist of multiple adjacent smallest subdomains L_0 ,

$$L = n L_0, \tag{2}$$

with $n \in \{1, 2, 4, 8, 16\}$ for LASSO and $n \in \{1, 2, 4, 8, 16, 32\}$ for RICO. This yields K subdomains of size L, with

$$K = \left(\frac{D_x}{L}\right)^2 = \left(\frac{D_y}{L}\right)^2.$$
 (3)

In each of these *K* subdomains, indexed by subscript *k*, the number of clouds *N* of size *l* are then counted, indicated as $N_k(L, l)$. In this process no cloud is counted twice; in other words, should a cloud object cross subdomain boundaries then the cloud is included in the count for the subdomain in which it is identified first. As a result, the sum of N_k over all subdomains always exactly equals the total cloud number counted for the full domain.



FIG. 5: The normalized cloud number density \mathcal{N}_* for the first RICO snapshot (t = 8hr) calculated for various subdomain sizes. The densities are plotted as a two-dimensional PDF in (l, \mathcal{N}_*) space. The shading indicates the probability of occurrence among the subdomains, normalized such that the most frequently occuring value of \mathcal{N}_* at a given value of l has a value of 1 (red). The curved dotted line is the least squares fit of a powerlaw-exponential function to the full domain CSD, while the horizontal dotted line indicates the value of a single cloud in the average subdomain, calculated as $(\sum_l \overline{N}(L,l)dl)^{-1}$.

With N_k known this finally allows calculating the mean cloud number per subdomain, $\overline{N}(L, l)$, as well as the associated standard deviation $\sigma(L, l)$,

$$\overline{N}(L,l) = \frac{1}{K} \sum_{k=1}^{K} N_k(L,l)$$
(4)

$$\sigma(L,l) = \left(\frac{1}{K}\sum_{k=1}^{K} \left[N_k(L,l) - \overline{N}(L,l)\right]^2\right)^{\frac{1}{2}}$$
(5)

Equation (5) expresses how variable the number of clouds of a certain size *l* is when considering subdomains of size *L*. The variable σ is thus a measure of the CSD-internal variability, visible in Fig. 4 as the vertical spread in the right tail of the CSD. In practice, it is only interpreted for $K \ge 4$ to ensure statistical significance. It is important to note that this definition of σ accounts for the impact of no clouds occurring in a subdomain (i.e. $N_k = 0$). This has the potential to significantly increase σ in situations when the subdomain size is of the same order of magnitude as the cloud spacing, as will be illustrated later.

3. Results

a. Number densities in subdomains

Figure 5 shows a subdomain analysis of the size density of the number of clouds (the "number density") of the first 3D snapshot from the RICO case. By visual comparison in Fig. 1 this cloud field shows the lowest degree of spatial organization, so that the CSD-internal varibility will be predominantly due to subsampling. Unlike the CSDs as shown in Fig. 4 the data are now plotted not as histograms but as a two-dimensional PDFs in (l, \mathcal{N}_*) -space. The normalized size density for subdomains $\mathcal{N}_*(L,l)$ is obtained by dividing N(L,l) by the average total number of clouds in a subdomain of size L,

$$\mathscr{N}_{*}(L,l) = \frac{N(L,l)}{\sum_{l} \overline{N}(L,l)} \, \mathrm{d}l^{-1},\tag{6}$$

Normalized number densities are commonly used in studies of cumulus cloud population, an advantage being that per definition the integral of \mathcal{N}_* with l always adds up to 1. As a result, in (l, \mathcal{N}_*) -space the vertical position of the density is preserved for subdomain size, which facilitates their comparison. The two-dimensional PDF is created by sorting the number densities of all subdomains onto a 100x100 histogram in the log-log range shown. The shading reflects the relative probability of occurrence of a value of \mathcal{N}_* among the subdomains for a given value of *l*, with red marking the value that occurs most. Two lines are added for reference; a powerlaw-exponential fit on the full domain CSD as proposed by Ding et al. (2014), and a horizontal line indicating the value of a single cloud occurrence in the average subdomain. As a result of the use of the normalized number density on the y-axis, the single cloud value increases with smaller L due to the decreasing number of clouds present in the subdomain.

The first panel in Fig. 5 shows the analysis for n = 32, so that the subdomain equals the full domain. As a result, the plotted data is equivalent to that shown in Fig. 4a. A few aspects catch the eye. First, the powerlaw-exponental fit well captures the shape of the CSD in this case. Second, the horizontal spacing in the left tail of the PDF reflects the horizontal discretization of the LES, allowing only a few specific cloud sizes. Third, the vertical spread in the data increases towards the largest sizes, which is an expression of an increase in the CSD-internal variability. Also, the data tends to organize along horizontal lines, with the lowest line of points (with lowest \mathcal{N}_* value) coinciding with the single-cloud line. This implies that only one cloud object of this size is present in the domain.

The following panels show the analysis for decreasing subdomain size *L*, for $n \in \{16, 8, 4, 2, 1\}$. With decreasing



FIG. 6: The probability P that no clouds are present in a subdomain, as a function of cloud size l, for a range of subdomain sizes L.

L the single cloud line starts to intersect the full domain fit at ever smaller *l*. To the right of the intersection size the points in the PDF are situated at the single cloud line, which thus again acts as a minimum possible value of \mathcal{N}_* . For the smallest domain size L = 1.6 km only N = 2, N = 3and N = 4 occur. This behavior can be explained by the subdomain becoming comparable or even smaller than the cloud spacing, so that sometimes no clouds are present in it. This binary-like occurrence of clouds can significantly boost the variance σ^2 at these sizes. More insight is provided by Fig. 6, showing the probability to encounter a cloud-free subdomain. In general, this probability is highest for the largest clouds and lowest for the smallest, a result which supports this hypothesis. This transition shifts with decreasing subdomain size L, with no-cloud occurrences becoming more frequent at smaller l. At smallto-intermediate L the transition is gradual, while for large L the transition is abrupt, almost binary. From these results one concludes that the no-cloud occurrence and its impact on the internal variability of the CSD is in principle a stochastic behavior, reflecting the subsampling of populations.

b. Stochastic variability

The next step is to quantify the internal variability of the CSD due to subsampling as analyzed in the previous section. This is achieved by calculating the variance in cloud number σ^2 for all cloud- and subdomain sizes, as defined by (5). To make sure that the results will only reflect subsampling, again the first RICO timepoint (t=8hr) is used that shows the least signs of spatial organization. Figure 7a shows σ normalized by the average cloud number per subdomain \overline{N} as a function of subdomain size *L*. Data points are equidistant in the horizontal due to the set of *n*-values that was chosen. For all cloud sizes a well-defined monotone relation exists between these variables in log-log space, indicating a strong powerlaw scaling. Variabil-



FIG. 7: Log-log plot of the normalized standard deviation in cloud number for the first snapshot (t = 8hr) of the RICO case, plotted as a function of a) subdomain size L and b) subdomain size divided by cloud size L/l. A selection of sizes across the spectrum is shown, as indicated by the colors. The dotted line represents a fit of the baseline model for the impact of subsampling on cloud number variability with b = -1, as discussed in Section 3c.

ity in cloud number also increases with cloud size, shifting the relation vertically in a fairly equidistant way, leaving the slope unaffected. In general the relation is best defined at the smaller domain sizes, but shows more noise at the larger subdomain sizes. This is caused by the number of subdomains *K* as used in the variance calculation decreasing quadratically with domain size (e.g. K = 4 at L = 0.5D).

The fact that the powerlaw scaling is similar for all cloud sizes, and that the lines of constant cloud size are more or less equidistant in the vertical, motivates reformulation in terms of two dimensionless groups. Figure 7b shows the same log-log plot but now with subdomain size *L* divided by cloud size *l*. This transformation yields a data collapse. A first order estimation of the slope suggests a powerlaw exponent b = -1, a fit of which indeed captures the dependence to a high degree (dotted line). This dependence reflects an inverse linear relation between the dimensionless groups (σ/\overline{N}) and (L/l).

c. The impact of subsampling

What could explain the dependence of the CSD variability on both the subdomain size L and the cloud size l? Let us first consider the inverse linear dependence on L. One expects the variability in the cloud number present in a subdomain N to be significant when L is of the same order of magnitude as the spacing between the clouds. This can be understood from considering the hypothetical case of a population of clouds of a single size l which is regularly distributed horizontally, with each cloud and



FIG. 8: Schematic illustration of the subsampling of an equally spaced single-size cloud population. The red box indicates the subdomain of size L, while the blue gridboxes represent the spatial unit occupied by a single cloud and its spacing l_s . Variables are explained in the text.



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FIG. 9: a) Same as Fig. 7a but now showing only a single cloud size l = 25 m as a function of time in the RICO case. b) Powerlaw exponent *b* as a function of cloud size and time in the RICO case, resulting from linear fits in $\log(\sigma/\overline{N})$, $\log(L/l)$ space.

the cloud-free area around it occupying a "unit area" of spacing size l_s . This situation is illustrated in Fig. 8. A subdomain is considered that covers a number of cloud-containing units, some of which partially (illustrated here as one row and one column at the edge). In this situation the number of clouds in subdomain N_k can be written as

$$N_k = \left(\frac{L}{l_s}\right)^2 \pm \mathscr{O}\left(\frac{L}{l_s}\right). \tag{7}$$

The first term on the right-hand-side is a quadratic function of L, and represents the number of clouds included in the subdomain that one expects statistically after many subdomains are considered, and is conceptually equal to \overline{N} . In contrast, the second term is a *linear* function of L, and represents the number of clouds near the subdomain's edges that may or may not be situated inside its area. For the schematic cloud scene shown in Fig. 8 with $L = 3.5l_s$ this gives $N_k = 12.25 \pm 3.5$, implying that the number of clouds in a single arbitrary snapshot can be anywhere between 9 and 16. The larger L/l_s , the smaller the remainder term relative to the quadratic term, and the smaller the variation in the estimate of N_k . This corresponds to a less subsampled cloud population. What this implies for the cloud number variance σ^2 can then simply be understood by substituting (7) in (5), which gives

$$\boldsymbol{\sigma} = \mathscr{O}\left(\frac{L}{l_s}\right),\tag{8}$$

where the quadratic term has disappeared because of the appearence of \overline{N} in the definition of variance, so that only

the linear "boundary effect" term remains. Finally normalizing by the average cloud number \overline{N} yields

$$\frac{\sigma}{\overline{N}} \propto \left(\frac{L}{l_s}\right) \left(\frac{L}{l_s}\right)^{-2} = \left(\frac{L}{l_s}\right)^{-1}.$$
(9)

This proportionality explains the inverse linear dependence of the normalized standard deviation in cloud number on the subdomain size that is purely due to subsampling. The powerlaw-exponent of -1 can thus be considered a baseline model for this process.

What is also clear is that the effect of subsampling is first experienced by larger clouds. We speculate that this is explained by the observation that larger clouds typically having larger spacing l_s (Joseph and Cahalan 1990). The relation between l and l_s could carry some dependence on the cumulus regime, but also critically depends on how the cloud size l and the neighbor spacing l_s are exactly defined. More research for more cumulus cases is needed to gain insight.

d. The impact of spatial organization

The analysis is now repeated for all six RICO fields, which become increasingly organized as time progresses, as shown in Fig. 1. The time-evolution of the relation between σ/\overline{N} and *L* is shown in Figure 9a, first for a single cloud size (l = 25m). In general the variability increases with time, already suggesting that some process is affecting the occurrence of the smallest clouds on the grid. Most importantly, the increase in variability is not uniform, but



FIG. 10: Same as Fig. 9b but now for the LASSO 20160518, 20160619, 20160625, 20160716 and 20160818 cases. a) - e) Each case individually, and f) composite average of all five cases. Non-simulated time-periods are shaded light-grey.

its amplitude increases with subdomain size *L*. Powerlaw scaling is still evident at all timepoints, but with an increasingly reduced exponent. This means that, especially at larger *L*, the variance in the number of small clouds is larger than can be expected from pure subsampling effects, as expressed by the b = -1 baseline model.

More insight is provided by Fig. 9b, showing the powerlaw exponent for all cloud sizes l as a function of time.

Two features stand out. First, the maximum cloud size in the domain increases with time, approaching 10 km at the $t = \{24, 32\}$ hr timepoints. This reflects the emergence of large cloud structures as visible in Fig. 1, although their number is still very small. Second, the impact of the spatial organization in the population on the powerlaw exponent appears limited to the smaller end of the distribution, and gradually disappears above about l = 250m. These



FIG. 11: Same as Fig. 7b but now including all snapshots from all 6 cases. All clouds with sizes l > 250m are included. The dashed line is a least-squares fit of the function $F = F_0 + b\log(L/l)$ to the composite data, yielding b = -0.92 at $F_0 = 1.2$. For reference, the fit of the baseline model with prescribed b = -1, as already shown in Fig. 7b, is also included (dotted line).

two findings suggest that spatial organization apparently affects both ends of the CSD. On the one end the size of the largest cloud increases, while on the other the variability in the number of small clouds is enhanced. This reflects that small clouds "live" on top of larger structures (e.g. cold pools), favoring or inhibiting their formation.

To investigate the robustness of the impact of spatial organization on the CSD variability the analysis is now repeated for all LASSO cases, as shown in Fig. 10a-e. The time-evolution of the CSD is similar in all LASSO cases, with a gradually increasing maximum cloud size in the period after cloud onset A reduction in the powerlaw exponent at the smaller cloud sizes is also present, however it is only visible towards the end of the day. A clear difference with the RICO case is the much shorter time-span during which this happens. This suggests that the physical process of spatial organization causing this behavior is of a slightly different nature. After inspecting the cloud mask fields in Fig. 2 one notices that towards the end of the day small clouds start to surround the bigger clouds in the population. One speculates that this is caused by big clouds falling apart, towards sunset when turbulence dies. In that sense one expects that this behavior is a typical and prototype feature of diurnal cycles of shallow cumulus over land. This motivates compositing the five cases to enhance the statistical significance, as shown in Fig. 10f. The compositing indeed makes the contoured field smoother. In addition, the compositing emphasizes that the reduction of the powerlaw exponent in the last hours of convective cloud existence is a robust feature.

4. Discussion

The results obtained so far for the RICO and LASSO cases suggest that the impact of spatial organization on the variability in cloud number is superimposed on the impact of subsampling alone. Two different mechanisms of spatial organization have been encountered in this study, including cold pool dynamics in the RICO case and cloud decay in the LASSO cases. But apparently the impact of subsampling always takes place in the same way. This is further illustrated by Fig. 11, showing the scaled variances of all timepoints from all cases in one frame. All cloud sizes l > 250 m are included, thus excluding the size range that is most affected by organization in the cases considered. Combining all fields from all cases still yields a data collapse for this size-range. A powerlaw fit is included in Fig. 11 (dashed line), yielding b = -0.92 which is very close to the baseline model of b = -1 as formulated in

Section c (dotted line). This result suggests that the baseline model for the impact of subsampling on the variability in the cumulus CSD is indeed generally applicable.

This motivates expressing the CSD-internal variability in terms of two dimensionless groups,

$$\left(\frac{\sigma}{\overline{N}}\right) = a \left(\frac{L}{l}\right)^b \tag{10}$$

with exponent b written as

$$b = -1 + \mathscr{F}(\operatorname{Org}). \tag{11}$$

The baseline value of b = -1 exclusively reflects the impact of subsampling due to a too small subdomain size. On top of this baseline exponent the impact of the degree of spatial organization in the population is then superimposed, through an extra term depending on the degree of organization. While the latter dependence still needs to be established, one option could be the Org parameter as proposed by Mapes and Neale (2011). Parameter *a*, used here as constant of proportionality to fit scaling relation (10) to the LES data, effectively translates the relation vertically in loglog space. A shift in this parameter thus reflects an overall change in the variability for all subdomain sizes; if this carries an additional dependence on Org is still unclear, and requires more research.

The scaling relation (10) has various potential uses. For example, it can inform the development of parameterizations in the grey zone, by acting as a benchmark relation that stochastic parameterizations of cumulus cloud populations have to reproduce. This is in particularly applicable to schemes that are formulated in terms of reconstructed CSDs. Complex population dynamics models might let the number density grow from interaction between sizes, while simpler ones using assumed functional forms for the CSD could have the impact of subsampling and spatial organization as established in this study superimposed. Regardless of the design of the stochastic parameterization, they can be called succesful when they reproduce the b = -1 scaling for unorganized cloud populations in the grey zone of convection.

The results obtained in this study highlight the important role of cloud spacing in controlling the stochastic variability. Normalizing subdomain size by cloud size yields a data collapse in the scaled variance, argued here to reflect that cloud size is proportional to the cloud spacing. The data collapse suggests that the relation between cloud size and cloud spacing could be pretty robust. More research is needed to confirm and quantify this proportionality. Previous observational studies have indeed hinted at its existence in nature (e.g. Joseph and Cahalan 1990; Sengupta et al. 1990), but typically these relations still carried considerable scatter. In addition, the definition of cloud spacing used in most previous studies is not equivalent to what is used in this study, which is the spacing between clouds *of equal size*. Exploring both observed and simulated datasets to investigate various definitions of cloud spacing is a future research topic.

This study also sheds more light on the nature of spatial organization in shallow cumulus cloud fields. Two types were encountered in the cases covered, including cold pool formation in the marine case and large cloud decay in the diurnal cycle cases. While this study does not explain how the organization takes place, what the results do reveal is that both types of spatial organization are associated with a similar shift in the powerlaw exponent in the relation between (σ/\overline{N}) and (L/l). Perhaps this shared behavior is something typical for spatial organization. But more cases need to be investigated before being able to claim general applicability of this behavior, for example for deeper precipitating convection.

Some aspects of our analysis can affect the scaling. The first is numerical and concerns the LES discretization, which could artificially affect the variability of the smallest clouds. This is investigated in more detail in the Appendix. We find thaat while the powerlaw scaling in the cloud number variability (slope) is unaffected, the amplitude increases at cloud sizes smaller than $4 \times \Delta x$. A close proximity to discretization thus slightly enhances the variability in clouds number. Note that these numerically affected clouds were not included in the fitting exercise shown in Fig. 11. Nevertheless, because the spatial organization affects the smallest clouds most, additional simulations at even higher resolution are still advisable to investigate at which point convergence takes place. A second important aspect concerns the definition of cloud size, for which more complex alternatives can be used. For example, cloud objects could be tracked through time, which might change how their variability behaves. Perhaps this study, based on a simple definition of size, can serve as a starting point for such analyses.

5. Summary, conclusions and outlook

The main results of this study can be summarized as follows:

- Multiple three-dimensional snapshots from various large-domain LES simulations of marine and continental shallow cumulus cases are used to investigate the behavior of variability in cumulus cloud number due to subsampling and spatial organization;
- Strong powerlaw scaling is found in the relation between cloud number variability and subdomain size, reflecting an inverse linear relation;
- Cloud spacing as a function of cloud size is a crucial parameter, explaining the data collapse across cases and timepoints;

- The impact of spatial organization on the variability in the CSD is found to act on top of this baseline model by enhancing its amplitude, in particular at the smaller cloud sizes. The powerlaw scaling is still preserved but with a reduced exponent. This impact reflects that the small clouds start to live on top of larger-scale structures, favoring or inhibiting their formation;
- Compositing all continental cumulus cases suggests the existence of a prototype diurnal time-dependence in the spatial organization, partially reflecting the decay of large convective cumulus clouds towards sunset;
- A simple expression for the CSD-internal variability is proposed that is formulated in terms of two dimensionless groups, and which captures the impact of both subsampling and spatial organization.

The subdomain analysis adopted in this study needs to be repeated for many more cumulus cloud scenes, in order to broaden the parameter space of environmental largescale conditions and thus enhance the statistical significance of the obtained results. This is required to better calibrate constants of proportionality, and also to better understand the nature of spatial organization in cumulus cloud populations. Of particular relevance is the behavior of nearest neighbor spacing between clouds of a certain size, as this dependence is at the foundation of the scaling relation found in this study. In an ongoing effort by the authors, data from super-large-domain LES experiments with the new Icosahedral Non-hydrostatic (ICON, Zängl et al. 2015) model in the Carribean dry season are used to this purpose. In addition, existing observational datasets of shallow cumulus cloud populations are revisited to seek observational support for the scaling found in the simulations. These data consist of satellite imagery as well as data from cloud-detecting ground-based imaging and remote sensing instrumentation at the ARM SGP site.

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FIG. A1: Same as Fig. 7b but now showing clouds with l < 250 m.

APPENDIX

Impacts of discretization

In Fig. A1 the variance scaling is investigated for cloud sizes smaller than $10 \times$ the LES discretization. For the clouds smaller than about $4 \times \Delta x$ the scaling relation starts to be situated above the baseline model (dotted line), but remains more or less parallel. This implies that at sizes near the discretization scale the CSD variability is somewhat larger than can be expected from subsampling alone. We speculate that this numerical artefact reflects on-off cloud behavior in single gridboxes that artificially boosts the variability in their number. However, most importantly, the powerlaw exponent is more or less unaffected by the proximity to discretization, meaning that the scaling is maintained even for barely resolved clouds.

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